

Package ‘spStack’

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Type Package

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Title Bayesian Geostatistics Using Predictive Stacking

Description Fits Bayesian hierarchical spatial process models for point-referenced Gaussian, Poisson, binomial, and binary data using stacking of predictive densities. It involves sampling from analytically available posterior distributions conditional upon some candidate values of the spatial process parameters and, subsequently assimilate inference from these individual posterior distributions using Bayesian predictive stacking. Our algorithm is highly parallelizable and hence, much faster than traditional Markov chain Monte Carlo algorithms while delivering competitive predictive performance. See Zhang, Tang, and Banerjee (2024) <[doi:10.48550/arXiv.2304.12414](https://doi.org/10.48550/arXiv.2304.12414)>, and, Pan, Zhang, Bradley, and Banerjee (2024) <[doi:10.48550/arXiv.2406.04655](https://doi.org/10.48550/arXiv.2406.04655)> for details.

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License GPL-3

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<https://span-18.github.io/spStack-dev/>

BugReports <https://github.com/SPan-18/spStack-dev/issues>

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spStack-package

spStack: Bayesian Geostatistics Using Predictive Stacking

Description

This package delivers functions to fit Bayesian hierarchical spatial process models for point-referenced Gaussian, Poisson, binomial, and binary data using stacking of predictive densities. It involves sampling from analytically available posterior distributions conditional upon some candidate values of the spatial process parameters for both Gaussian response model as well as non-Gaussian responses, and, subsequently assimilate inference from these individual posterior distributions using Bayesian predictive stacking. Our algorithm is highly parallelizable and hence, much faster than traditional Markov chain Monte Carlo algorithms while delivering competitive predictive performance.

In context of inference for spatial point-referenced data, Bayesian hierarchical models involve latent spatial processes characterized by spatial process parameters, which besides lacking substantive relevance in scientific contexts, are also weakly identified and hence, impedes convergence of MCMC algorithms. This motivates us to build methodology that involves fast sampling from posterior distributions conditioned on a grid of the weakly identified model parameters and combine the inference by stacking of predictive densities (Yao *et. al* 2018). We exploit the Bayesian conjugate linear modeling framework for the Gaussian case (Zhang, Tang and Banerjee 2024) and the generalized conjugate multivariate distribution theory (Pan, Zhang, Bradley and Banerjee 2024) to analytically derive the individual posterior distributions.

Details

Package: spStack
Type: Package
Version: 0.1.0
License: GPL-3

Accepts a formula, e.g., $y \sim x_1 + x_2$, for most regression models accompanied by candidate values of spatial process parameters, and returns posterior samples of the regression coefficients and the latent spatial random effects. Posterior inference or prediction of any quantity of interest proceed from these samples. Main functions are -

[spLMexact\(\)](#)
[spGLMexact\(\)](#)
[spLMstack\(\)](#)
[spGLMstack\(\)](#)

Author(s)

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References

Zhang L, Tang W, Banerjee S (2024). "Bayesian Geostatistics Using Predictive Stacking." [doi:10.48550/arXiv.2304.12414](https://doi.org/10.48550/arXiv.2304.12414).

Pan S, Zhang L, Bradley JR, Banerjee S (2024). "Bayesian Inference for Spatial-temporal Non-Gaussian Data Using Predictive Stacking." [doi:10.48550/arXiv.2406.04655](https://doi.org/10.48550/arXiv.2406.04655).

Yao Y, Vehtari A, Simpson D, Gelman A (2018). "Using Stacking to Average Bayesian Predictive Distributions (with Discussion)." *Bayesian Analysis*, **13**(3), 917-1007. [doi:10.1214/17BA1091](https://doi.org/10.1214/17BA1091).

See Also

Useful links:

- <https://github.com/SPan-18/spStack-dev>
- <https://span-18.github.io/spStack-dev/>
- Report bugs at <https://github.com/SPan-18/spStack-dev/issues>

cholUpdate *Different Cholesky factor updates*

Description

Provides functions that implements different types of updates of a Cholesky factor that includes rank-one update, single row/column deletion update and a block deletion update.

Usage

```
cholUpdateRankOne(A, v, alpha, beta, lower = TRUE)
```

```
cholUpdateDel(A, del.index, lower = TRUE)
```

```
cholUpdateDelBlock(A, del.start, del.end, lower = TRUE)
```

Arguments

A	an $n \times n$ triangular matrix
v	an $n \times 1$ matrix/vector
alpha	scalar; if not supplied, default is 1
beta	scalar; if not supplied, default is 1
lower	logical; if A is lower-triangular or not
del.index	an integer from 1 to n indicating the row/column to be deleted
del.start	an integer from 1 to n indicating the first row/column of a block to be deleted, must be at least 1 less than del.end
del.end	an integer from 1 to n indicating the last row/column of a block to be deleted, must be at least 1 more than del.start

Details

Suppose $B = AA^\top$ is a $n \times n$ matrix with A being its lower-triangular Cholesky factor. Then rank-one update corresponds to finding the Cholesky factor of the matrix $C = \alpha B + \beta vv^\top$ for some $\alpha, \beta \in \mathbb{R}$ given A (see, Krause and Igel 2015). Similarly, single row/column deletion update corresponds to finding the Cholesky factor of the $(n - 1) \times (n - 1)$ matrix B_i which is obtained by removing the i -th row and column of B , given A for some $i = 1, \dots, n$. Lastly, block deletion corresponds to finding the Cholesky factor of the $(n - n_k) \times (n - n_k)$ matrix B_I for a subset I of $\{1, \dots, n\}$ containing n_k consecutive indices, given the factor A .

Value

An $m \times m$ lower-triangular matrix with $m = n$ in case of cholUpdateRankOne(), $m = n - 1$ in case of cholUpdateDel(), and, $m = n - n_k$ in case of cholUpdateDelBlock() where n_k is the size of the block removed.

Author(s)

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References

Oswin Krause and Christian Igel. 2015. "A More Efficient Rank-one Covariance Matrix Update for Evolution Strategies". In *Proceedings of the 2015 ACM Conference on Foundations of Genetic Algorithms XIII (FOGA '15)*. Association for Computing Machinery, New York, NY, USA, 129-136. doi:[10.1145/2725494.2725496](https://doi.org/10.1145/2725494.2725496).

Examples

```
n <- 10
A <- matrix(rnorm(n^2), n, n)
A <- crossprod(A)
cholA <- chol(A)

## Rank-1 update
v <- 1:n
APlusvT <- A + tcrossprod(v)
cholA1 <- t(chol(APlusvT))
cholA2 <- cholUpdateRankOne(cholA, v, lower = FALSE)
print(all(abs(cholA1 - cholA2) < 1E-9))

## Single Row-deletion update
ind <- 2
A1 <- A[-ind, -ind]
cholA1 <- t(chol(A1))
cholA2 <- cholUpdateDel(cholA, del.index = ind, lower = FALSE)
print(all(abs(cholA1 - cholA2) < 1E-9))

## Block-deletion update
start_ind <- 2
end_ind <- 6
del_ind <- c(start_ind:end_ind)
A1 <- A[-del_ind, -del_ind]
cholA1 <- t(chol(A1))
cholA2 <- cholUpdateDelBlock(cholA, start_ind, end_ind, lower = FALSE)
print(all(abs(cholA1 - cholA2) < 1E-9))
```

get_stacking_weights *Optimal stacking weights*

Description

Obtains optimal stacking weights given leave-one-out predictive densities for each candidate model.

Usage

```
get_stacking_weights(log_loopd, solver = "ECOS")
```

Arguments

`log_loopd` an $n \times M$ matrix with i -th row containing the leave-one-out predictive densities for the i -th data point for the M candidate models.

`solver` specifies the solver to use for obtaining optimal weights. Default is "ECOS". Internally calls `CVXR::psolve()`.

Value

A list of length 2.

`weights` optimal stacking weights as a numeric vector of length M

`status` solver status, returns "optimal" if solver succeeded.

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References

Yao Y, Vehtari A, Simpson D, Gelman A (2018). "Using Stacking to Average Bayesian Predictive Distributions (with Discussion)." *Bayesian Analysis*, **13**(3), 917-1007. doi:[10.1214/17BA1091](https://doi.org/10.1214/17BA1091).

See Also

[CVXR::psolve\(\)](#), [sLMstack\(\)](#), [spGLMstack\(\)](#)

Examples

```
data(simGaussian)
dat <- simGaussian[1:100, ]

mod1 <- sLMstack(y ~ x1, data = dat,
  coords = as.matrix(dat[, c("s1", "s2")]),
  cor.fn = "matern",
  params.list = list(phi = c(1.5, 3),
    nu = c(0.5, 1),
    noise_sp_ratio = c(1)),
  n.samples = 1000, loopd.method = "exact",
  parallel = FALSE, solver = "ECOS", verbose = TRUE)

loopd_mat <- do.call('cbind', mod1$loopd)
w_hat <- get_stacking_weights(loopd_mat)
print(round(w_hat$weights, 4))
print(w_hat$status)
```

iDist	<i>Calculate distance matrix</i>
-------	----------------------------------

Description

Computes the inter-site Euclidean distance matrix for one or two sets of points.

Usage

```
iDist(coords.1, coords.2, ...)
```

Arguments

coords.1	an $n \times p$ matrix with each row corresponding to a point in p -dimensional space.
coords.2	an $m \times p$ matrix with each row corresponding to a point in p dimensional space. If this is missing then coords.1 is used.
...	currently no additional arguments.

Value

The $n \times n$ or $n \times m$ inter-site Euclidean distance matrix.

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Examples

```
n <- 10  
p1 <- cbind(runif(n),runif(n))  
m <- 5  
p2 <- cbind(runif(m),runif(m))  
D <- iDist(p1, p2)
```

simBinary	<i>Synthetic point-referenced binary data</i>
-----------	---

Description

Dataset of size 500, with a binary response variable indexed by spatial coordinates sampled uniformly from the unit square. The model includes one covariate and spatial random effects induced by a Matérn covariogram.

Usage

```
data(simBinary)
```

Format

a data.frame object.

s1, s2 2-D coordinates; latitude and longitude.

x1 a covariate sampled from the standard normal distribution.

y response vector (0/1).

z_true true spatial random effects that generated the data.

Details

With $n = 500$, the binary data is simulated using

$$y(s_i) \sim \text{Bernoulli}(\pi(s_i)), i = 1, \dots, n,$$

$$\pi(s_i) = \text{ilogit}(x(s_i)^\top \beta + z(s_i))$$

where the function `ilogit` refers to the inverse-logit function, the spatial effects $z \sim N(0, \sigma^2 R)$ with R being a $n \times n$ correlation matrix given by the Matérn covariogram

$$R(s, s') = \frac{(\phi |s - s'|)^\nu}{\Gamma(\nu) 2^{\nu-1}} K_\nu(\phi |s - s'|),$$

where ϕ is the spatial decay parameter and ν the spatial smoothness parameter. We have sampled the data with $\beta = (0.5, -0.5)$, $\phi = 5$, $\nu = 0.5$, and $\sigma^2 = 0.4$. This data can be generated with the code as given in the example below.

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See Also

[simGaussian](#), [simPoisson](#), [simBinom](#)

Examples

```
set.seed(1729)
n <- 500
beta <- c(0.5, -0.5)
phi0 <- 5
nu0 <- 0.5
spParams <- c(phi0, nu0)
spvar <- 0.4
sim1 <- sim_spData(n = n, beta = beta, cor.fn = "matern",
                  spParams = spParams, spvar = spvar, deltasq = deltasq,
                  family = "binary")
plot1 <- surfaceplot(sim1, coords_name = c("s1", "s2"), var_name = "z_true")

library(ggplot2)
plot2 <- ggplot(sim1, aes(x = s1, y = s2)) +
```



```

geom_point(aes(color = factor(y)), alpha = 0.75) +
scale_color_manual(values = c("red", "blue"), labels = c("0", "1")) +
guides(alpha = 'none') +
theme_bw() +
theme(axis.ticks = element_line(linewidth = 0.25),
      panel.background = element_blank(),
      panel.grid = element_blank(),
      legend.title = element_text(size = 10, hjust = 0.25),
      legend.box.just = "center", aspect.ratio = 1)

```

simBinom

Synthetic point-referenced binomial count data

Description

Dataset of size 500, with a binomial response variable indexed by spatial coordinates sampled uniformly from the unit square. The model includes one covariate and spatial random effects induced by a Matérn covariogram. The number of trials at each location is sampled from a Poisson distribution with mean 20.

Usage

```
data(simBinom)
```

Format

a data.frame object.

s1, s2 2-D coordinates; latitude and longitude.

x1 a covariate sampled from the standard normal distribution.

y response vector.

n_trials Number of trials at that location.

z_true true spatial random effects that generated the data.

Details

With $n = 500$, the count data is simulated using

$$y(s_i) \sim \text{Binomial}(m(s_i), \pi(s_i)), i = 1, \dots, n,$$

$$\pi(s_i) = \text{ilogit}(x(s_i)^\top \beta + z(s_i))$$

where the function `ilogit` refers to the inverse-logit function, the number of trials $m(s_i)$ is sampled from a Poisson distribution with mean 20, the spatial effects $z \sim N(0, \sigma^2 R)$ with R being a $n \times n$ correlation matrix given by the Matérn covariogram

$$R(s, s') = \frac{(\phi |s - s'|)^\nu}{\Gamma(\nu) 2^{\nu-1}} K_\nu(\phi |s - s'|),$$

where ϕ is the spatial decay parameter and ν the spatial smoothness parameter. We have sampled the data with $\beta = (0.5, -0.5)$, $\phi = 3$, $\nu = 0.5$, and $\sigma^2 = 0.4$. This data can be generated with the code as given in the example below.

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See Also

[simGaussian](#), [simPoisson](#), [simBinary](#)

Examples

```
set.seed(1729)
n <- 500
beta <- c(0.5, -0.5)
phi0 <- 3
nu0 <- 0.5
spParams <- c(phi0, nu0)
spvar <- 0.4
sim1 <- sim_spData(n = n, beta = beta, cor.fn = "matern",
                  spParams = spParams, spvar = spvar, deltasq = deltasq,
                  n_binom = rpois(n, 20),
                  family = "binomial")
plot1 <- surfaceplot(sim1, coords_name = c("s1", "s2"), var_name = "z_true")

library(ggplot2)
plot2 <- ggplot(sim1, aes(x = s1, y = s2)) +
  geom_point(aes(color = y), alpha = 0.75) +
  scale_color_distiller(palette = "RdYlGn", direction = -1,
                       label = function(x) sprintf("%.0f", x)) +
  guides(alpha = 'none') +
  theme_bw() +
  theme(axis.ticks = element_line(linewidth = 0.25),
        panel.background = element_blank(),
        panel.grid = element_blank(),
        legend.title = element_text(size = 10, hjust = 0.25),
        legend.box.just = "center", aspect.ratio = 1)
```

simGaussian

Synthetic point-referenced Gaussian data

Description

Dataset of size 500 with a Gaussian response variable, simulated with spatial coordinates sampled uniformly from the unit square. The model includes one covariate and spatial random effects induced by a Matérn covariogram.

Usage

```
data(simGaussian)
```

Format

a data.frame object.

s1, s2 2-D coordinates; latitude and longitude.

x1 a covariate sampled from the standard normal distribution.

y response vector.

z_true true spatial random effects that generated the data.

Details

The data is generated using the model

$$y = X\beta + z + \epsilon,$$

where the spatial effects $z \sim N(0, \sigma^2 R)$ is independent of the measurement error $\epsilon \sim N(0, \delta^2 \sigma^2 I_n)$ with δ^2 being the noise-to-spatial variance ratio and R being a $n \times n$ correlation matrix given by the Matérn covariogram

$$R(s, s') = \frac{(\phi|s - s'|)^\nu}{\Gamma(\nu)2^{\nu-1}} K_\nu(\phi|s - s'|),$$

where ϕ is the spatial decay parameter and ν the spatial smoothness parameter. We have sampled the data with $\beta = (2, 5)$, $\phi = 2$, $\nu = 0.5$, $\delta^2 = 1$ and $\sigma^2 = 0.4$. This data can be generated with the code as given in the example.

Author(s)

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See Also

[simPoisson](#), [simBinom](#), [simBinary](#)

Examples

```
set.seed(1729)
n <- 500
beta <- c(2, 5)
phi0 <- 2
nu0 <- 0.5
spParams <- c(phi0, nu0)
spvar <- 0.4
deltasq <- 1
sim1 <- sim_spData(n = n, beta = beta, cor.fn = "matern",
                  spParams = spParams, spvar = spvar, deltasq = deltasq,
                  family = "gaussian")
plot1 <- surfaceplot(sim1, coords_name = c("s1", "s2"), var_name = "z_true",
                    mark_points = TRUE)

plot1

library(ggplot2)
plot2 <- ggplot(sim1, aes(x = s1, y = s2)) +
```

```

geom_point(aes(color = y), alpha = 0.75) +
scale_color_distiller(palette = "RdYlGn", direction = -1,
                      label = function(x) sprintf("%.0f", x)) +
guides(alpha = 'none') + theme_bw() +
theme(axis.ticks = element_line(linewidth = 0.25),
      panel.background = element_blank(), panel.grid = element_blank(),
      legend.title = element_text(size = 10, hjust = 0.25),
      legend.box.just = "center", aspect.ratio = 1)
plot2

```

simPoisson

Synthetic point-referenced Poisson count data

Description

Dataset of size 500, with a Poisson distributed response variable indexed by spatial coordinates sampled uniformly from the unit square. The model includes one covariate and spatial random effects induced by a Matérn covariogram.

Usage

```
data(simPoisson)
```

Format

a `data.frame` object.

`s1`, `s2` 2-D coordinates; latitude and longitude.

`x1` a covariate sampled from the standard normal distribution.

`y` response vector.

`z_true` true spatial random effects that generated the data.

Details

With $n = 500$, the count data is simulated using

$$y(s_i) \sim \text{Poisson}(\lambda(s_i)), i = 1, \dots, n,$$

$$\log \lambda(s_i) = x(s_i)^\top \beta + z(s_i)$$

where the spatial effects $z \sim N(0, \sigma^2 R)$ with R being a $n \times n$ correlation matrix given by the Matérn covariogram

$$R(s, s') = \frac{(\phi |s - s'|)^\nu}{\Gamma(\nu) 2^{\nu-1}} K_\nu(\phi |s - s'|),$$

where ϕ is the spatial decay parameter and ν the spatial smoothness parameter. We have sampled the data with $\beta = (2, -0.5)$, $\phi = 5$, $\nu = 0.5$, and $\sigma^2 = 0.4$. This data can be generated with the code as given in the example below.

See Also

[simGaussian](#), [simBinom](#), [simBinary](#)

Examples

```
set.seed(1729)
n <- 500
beta <- c(2, -0.5)
phi0 <- 5
nu0 <- 0.5
spParams <- c(phi0, nu0)
spvar <- 0.4
sim1 <- sim_spData(n = n, beta = beta, cor.fn = "matern",
                  spParams = spParams, spvar = spvar, deltasq = deltasq,
                  family = "poisson")

# Plot an interpolated spatial surface of the true random spatial effects
plot1 <- surfaceplot(sim1, coords_name = c("s1", "s2"), var_name = "z_true")

# Plot the simulated count data
library(ggplot2)
plot2 <- ggplot(sim1, aes(x = s1, y = s2)) +
  geom_point(aes(color = y), alpha = 0.75) +
  scale_color_distiller(palette = "RdYlGn", direction = -1,
                       label = function(x) sprintf("%.0f", x)) +
  guides(alpha = 'none') + theme_bw() +
  theme(axis.ticks = element_line(linewidth = 0.25),
        panel.background = element_blank(), panel.grid = element_blank(),
        legend.title = element_text(size = 10, hjust = 0.25),
        legend.box.just = "center", aspect.ratio = 1)
```

sim_spData

Simulate spatial data on unit square

Description

Generates synthetic spatial data of different types where the spatial co-ordinates are sampled uniformly on an unit square. Different types include point-referenced Gaussian, Poisson, binomial and binary data. The design includes an intercept and fixed covariates sampled from a standard normal distribution.

Usage

```
sim_spData(n, beta, cor.fn, spParams, spvar, deltasq, family, n_binom)
```

Arguments

n	sample size.
beta	a p -dimensional vector of fixed effects.
cor.fn	a quoted keyword that specifies the correlation function used to model the spatial dependence structure among the observations. Supported covariance model key words are: 'exponential' and 'matern'.
spParams	a numeric vector containing spatial process parameters - e.g., spatial decay and smoothness.
spvar	value of spatial variance parameter.
deltasq	value of noise-to-spatial variance ratio.
family	a character specifying the distribution of the response as a member of the exponential family. Valid inputs are 'gaussian', 'poisson', 'binary', and 'binomial'.
n_binom	necessary only when family = 'binomial'. Must be a vector of length n that will specify the number of trials for each observation. If it is of length 1, then that value is considered to be the common value for the number of trials for all n observations.

Value

a data.frame object containing the columns -

- s1, s2 2D-coordinates in unit square
- x1, x2, ... covariates, not including intercept
- y response
- n_trials present only when binomial data is generated
- z_true true spatial effects with which the data is generated

Author(s)

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 Sudipto Banerjee sudipto@ucla.edu

Examples

```
set.seed(1729)
n <- 10
beta <- c(2, 5)
phi0 <- 2
nu0 <- 0.5
spParams <- c(phi0, nu0)
spvar <- 0.4
deltasq <- 1
sim1 <- sim_spData(n = n, beta = beta, cor.fn = "matern",
                  spParams = spParams, spvar = spvar, deltasq = deltasq,
                  family = "gaussian")
```

spGLMexact

*Univariate Bayesian spatial generalized linear model***Description**

Fits a Bayesian spatial generalized linear model with fixed values of spatial process parameters and some auxiliary model parameters. The output contains posterior samples of the fixed effects, spatial random effects and, if required, finds leave-one-out predictive densities.

Usage

```
spGLMexact(
  formula,
  data = parent.frame(),
  family,
  coords,
  cor.fn,
  priors,
  spParams,
  boundary = 0.5,
  n.samples,
  loopd = FALSE,
  loopd.method = "exact",
  CV.K = 10,
  loopd.nMC = 500,
  verbose = TRUE,
  ...
)
```

Arguments

formula	a symbolic description of the regression model to be fit. See example below.
data	an optional data frame containing the variables in the model. If not found in data, the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>spGLMexact</code> is called.
family	Specifies the distribution of the response as a member of the exponential family. Supported options are 'poisson', 'binomial' and 'binary'.
coords	an $n \times 2$ matrix of the observation coordinates in \mathbb{R}^2 (e.g., easting and northing).
cor.fn	a quoted keyword that specifies the correlation function used to model the spatial dependence structure among the observations. Supported covariance model keywords are: 'exponential' and 'matern'. See below for details.
priors	(optional) a list with each tag corresponding to a hyperparameter name and containing hyperprior details. Valid tags include <code>V.beta</code> , <code>nu.beta</code> , <code>nu.z</code> and <code>sigmaSq.xi</code> . Values of <code>nu.beta</code> and <code>nu.z</code> must be at least 2.1. If not supplied, uses defaults.

spParams	fixed values of spatial process parameters.
boundary	Specifies the boundary adjustment parameter. Must be a real number between 0 and 1. Default is 0.5.
n.samples	number of posterior samples to be generated.
loopd	logical. If loopd=TRUE, returns leave-one-out predictive densities, using method as given by loopd.method. Default is FALSE.
loopd.method	character. Ignored if loopd=FALSE. If loopd=TRUE, valid inputs are 'exact', 'CV' and 'PSIS'. The option 'exact' corresponds to exact leave-one-out predictive densities which requires computation almost equivalent to fitting the model n times. The options 'CV' and 'PSIS' are faster and they implement K -fold cross validation and Pareto-smoothed importance sampling to find approximate leave-one-out predictive densities (Vehtari <i>et al.</i> 2017).
CV.K	An integer between 10 and 20. Considered only if loopd.method='CV'. Default is 10 (as recommended in Vehtari <i>et al.</i> 2017).
loopd.nMC	Number of Monte Carlo samples to be used to evaluate leave-one-out predictive densities when loopd.method is set to either 'exact' or 'CV'.
verbose	logical. If verbose = TRUE, prints model description.
...	currently no additional argument.

Details

With this function, we fit a Bayesian hierarchical spatial generalized linear model by sampling exactly from the joint posterior distribution utilizing the generalized conjugate multivariate distribution theory (Bradley and Clinch 2024). Suppose $\chi = (s_1, \dots, s_n)$ denotes the n spatial locations the response y is observed. Let $y(s)$ be the outcome at location s endowed with a probability law from the natural exponential family, which we denote by

$$y(s) \sim \text{EF}(x(s)^\top \beta + z(s); b, \psi)$$

for some positive parameter $b > 0$ and unit log partition function ψ . We consider the following response models based on the input supplied to the argument family.

'poisson' It considers point-referenced Poisson responses $y(s) \sim \text{Poisson}(e^{x(s)^\top \beta + z(s)})$. Here, $b = 1$ and $\psi(t) = e^t$.

'binomial' It considers point-referenced binomial counts $y(s) \sim \text{Binomial}(m(s), \pi(s))$ where, $m(s)$ denotes the total number of trials and probability of success $\pi(s) = \text{ilogit}(x(s)^\top \beta + z(s))$ at location s . Here, $b = m(s)$ and $\psi(t) = \log(1 + e^t)$.

'binary' It considers point-referenced binary data (0 or, 1) i.e., $y(s) \sim \text{Bernoulli}(\pi(s))$, where probability of success $\pi(s) = \text{ilogit}(x(s)^\top \beta + z(s))$ at location s . Here, $b = 1$ and $\psi(t) = \log(1 + e^t)$.

The hierarchical model is given as

$$\begin{aligned} y(s_i) \mid \beta, z, \xi &\sim \text{EF}(x(s_i)^\top \beta + z(s_i) + \xi_i - \mu_i; b_i, \psi_y), i = 1, \dots, n \\ \xi \mid \beta, z, \sigma_\xi^2, \alpha_\epsilon &\sim \text{GCM}_c(\dots), \\ \beta \mid \sigma_\beta^2 &\sim N(0, \sigma_\beta^2 V_\beta), \quad \sigma_\beta^2 \sim \text{IG}(\nu_\beta/2, \nu_\beta/2) \\ z \mid \sigma_z^2 &\sim N(0, \sigma_z^2 R(\chi; \phi, \nu)), \quad \sigma_z^2 \sim \text{IG}(\nu_z/2, \nu_z/2), \end{aligned}$$

where $\mu = (\mu_1, \dots, \mu_n)^\top$ denotes the discrepancy parameter. We fix the spatial process parameters ϕ and ν and the hyperparameters $V_\beta, \nu_\beta, \nu_z$ and σ_ξ^2 . The term ξ is known as the fine-scale variation term which is given a conditional generalized conjugate multivariate distribution as prior. For more details, see Pan *et al.* 2024. Default values for $V_\beta, \nu_\beta, \nu_z, \sigma_\xi^2$ are diagonal with each diagonal element 100, 2.1, 2.1 and 0.1 respectively.

Value

An object of class spGLMexact, which is a list with the following tags -

priors details of the priors used, containing the values of the boundary adjustment parameter (boundary), the variance parameter of the fine-scale variation term (simasq.xi) and others.

samples a list of length 3, containing posterior samples of fixed effects (beta), spatial effects (z) and the fine-scale variation term (xi).

loopd If loopd=TRUE, contains leave-one-out predictive densities.

model.params Values of the fixed parameters that includes phi (spatial decay), nu (spatial smoothness).

The return object might include additional data that can be used for subsequent prediction and/or model fit evaluation.

Author(s)

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References

Bradley JR, Clinch M (2024). "Generating Independent Replicates Directly from the Posterior Distribution for a Class of Spatial Hierarchical Models." *Journal of Computational and Graphical Statistics*, **0**(0), 1-17. doi:[10.1080/10618600.2024.2365728](https://doi.org/10.1080/10618600.2024.2365728).

Pan S, Zhang L, Bradley JR, Banerjee S (2024). "Bayesian Inference for Spatial-temporal Non-Gaussian Data Using Predictive Stacking." doi:[10.48550/arXiv.2406.04655](https://doi.org/10.48550/arXiv.2406.04655).

Vehtari A, Gelman A, Gabry J (2017). "Practical Bayesian Model Evaluation Using Leave-One-out Cross-Validation and WAIC." *Statistics and Computing*, **27**(5), 1413-1432. ISSN 0960-3174. doi:[10.1007/s1122201696964](https://doi.org/10.1007/s1122201696964).

See Also

[spLMexact\(\)](#)

Examples

```
# Example 1: Analyze spatial poisson count data
data(simPoisson)
dat <- simPoisson[1:10, ]
mod1 <- spGLMexact(y ~ x1, data = dat, family = "poisson",
                  coords = as.matrix(dat[, c("s1", "s2")]),
                  cor.fn = "matern",
                  spParams = list(phi = 4, nu = 0.4),
```

```

n.samples = 100, verbose = TRUE)

# summarize posterior samples
post_beta <- mod1$samples$beta
print(t(apply(post_beta, 1, function(x) quantile(x, c(0.025, 0.5, 0.975))))))

# Example 2: Analyze spatial binomial count data
data(simBinom)
dat <- simBinom[1:10, ]
mod2 <- spGLMexact(cbind(y, n_trials) ~ x1, data = dat, family = "binomial",
  coords = as.matrix(dat[, c("s1", "s2")]),
  cor.fn = "matern",
  spParams = list(phi = 3, nu = 0.4),
  n.samples = 100, verbose = TRUE)

# summarize posterior samples
post_beta <- mod2$samples$beta
print(t(apply(post_beta, 1, function(x) quantile(x, c(0.025, 0.5, 0.975))))))

# Example 3: Analyze spatial binary data
data(simBinary)
dat <- simBinary[1:10, ]
mod3 <- spGLMexact(y ~ x1, data = dat, family = "binary",
  coords = as.matrix(dat[, c("s1", "s2")]),
  cor.fn = "matern",
  spParams = list(phi = 4, nu = 0.4),
  n.samples = 100, verbose = TRUE)

# summarize posterior samples
post_beta <- mod3$samples$beta
print(t(apply(post_beta, 1, function(x) quantile(x, c(0.025, 0.5, 0.975))))))

```

spGLMstack

Bayesian spatial generalized linear model using predictive stacking

Description

Fits Bayesian spatial generalized linear model on a collection of candidate models constructed based on some candidate values of some model parameters specified by the user and subsequently combines inference by stacking predictive densities. See Pan, Zhang, Bradley, and Banerjee (2024) for more details.

Usage

```

spGLMstack(
  formula,
  data = parent.frame(),
  family,
  coords,

```

```

    cor.fn,
    priors,
    params.list,
    n.samples,
    loopd.controls,
    parallel = FALSE,
    solver = "ECOS",
    verbose = TRUE,
    ...
)

```

Arguments

formula	a symbolic description of the regression model to be fit. See example below.
data	an optional data frame containing the variables in the model. If not found in data, the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>spLMstack</code> is called.
family	Specifies the distribution of the response as a member of the exponential family. Supported options are 'poisson', 'binomial' and 'binary'.
coords	an $n \times 2$ matrix of the observation coordinates in \mathbb{R}^2 (e.g., easting and northing).
cor.fn	a quoted keyword that specifies the correlation function used to model the spatial dependence structure among the observations. Supported covariance model keywords are: 'exponential' and 'matern'. See below for details.
priors	(optional) a list with each tag corresponding to a parameter name and containing prior details. Valid tags include <code>V.beta</code> , <code>nu.beta</code> , <code>nu.z</code> and <code>sigmaSq.xi</code> .
params.list	a list containing candidate values of spatial process parameters for the <code>cor.fn</code> used, and, the boundary parameter.
n.samples	number of posterior samples to be generated.
loopd.controls	a list with details on how leave-one-out predictive densities (LOO-PD) are to be calculated. Valid tags include <code>method</code> , <code>CV.K</code> and <code>nMC</code> . The tag <code>method</code> can be either 'exact' or 'CV'. If sample size is more than 100, then the default is 'CV' with <code>CV.K</code> equal to its default value 10 (Gelman <i>et al.</i> 2024). The tag <code>nMC</code> decides how many Monte Carlo samples will be used to evaluate the leave-one-out predictive densities, which must be at least 500 (default).
parallel	logical. If <code>parallel=FALSE</code> , the parallelization plan, if set up by the user, is ignored. If <code>parallel=TRUE</code> , the function inherits the parallelization plan that is set by the user via the function <code>future::plan()</code> only. Depending on the parallel backend available, users may choose their own plan. More details are available at https://cran.R-project.org/package=future .
solver	(optional) Specifies the name of the solver that will be used to obtain optimal stacking weights for each candidate model. Default is 'ECOS'. Users can use other solvers supported by the <code>CVXR-package</code> package.
verbose	logical. If TRUE, prints model-specific optimal stacking weights.
...	currently no additional argument.

Details

Instead of assigning a prior on the process parameters ϕ and ν , the boundary adjustment parameter ϵ , we consider a set of candidate models based on some candidate values of these parameters supplied by the user. Suppose the set of candidate models is $\mathcal{M} = \{M_1, \dots, M_G\}$. Then for each $g = 1, \dots, G$, we sample from the posterior distribution $p(\sigma^2, \beta, z \mid y, M_g)$ under the model M_g and find leave-one-out predictive densities $p(y_i \mid y_{-i}, M_g)$. Then we solve the optimization problem

$$\begin{aligned} \max_{w_1, \dots, w_G} \quad & \frac{1}{n} \sum_{i=1}^n \log \sum_{g=1}^G w_g p(y_i \mid y_{-i}, M_g) \\ \text{subject to} \quad & w_g \geq 0, \sum_{g=1}^G w_g = 1 \end{aligned}$$

to find the optimal stacking weights $\hat{w}_1, \dots, \hat{w}_G$.

Value

An object of class `spGLMstack`, which is a list including the following tags -

`family` the distribution of the responses as indicated in the function call

`samples` a list of length equal to total number of candidate models with each entry corresponding to a list of length 3, containing posterior samples of fixed effects (beta), spatial effects (z) and fine-scale variation term (xi) for that particular model.

`lopd` a list of length equal to total number of candidate models with each entry containing leave-one-out predictive densities under that particular model.

`lopd.method` a list containing details of the algorithm used for calculation of leave-one-out predictive densities.

`n.models` number of candidate models that are fit.

`candidate.models` a matrix with `n_model` rows with each row containing details of the model parameters and its optimal weight.

`stacking.weights` a numeric vector of length equal to the number of candidate models storing the optimal stacking weights.

`run.time` a `proc_time` object with runtime details.

`solver.status` solver status as returned by the optimization routine.

The return object might include additional data that is useful for subsequent prediction, model fit evaluation and other utilities.

Author(s)

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References

Pan S, Zhang L, Bradley JR, Banerjee S (2024). "Bayesian Inference for Spatial-temporal Non-Gaussian Data Using Predictive Stacking." doi:10.48550/arXiv.2406.04655.

Vehtari A, Simpson D, Gelman A, Yao Y, Gabry J (2024). "Pareto Smoothed Importance Sampling." *Journal of Machine Learning Research*, **25**(72), 1-58. URL <https://jmlr.org/papers/v25/19-556.html>.

See Also

[spGLMexact\(\)](#), [spLMstack\(\)](#)

Examples

```
data("simPoisson")
dat <- simPoisson[1:100,]
mod1 <- spGLMstack(y ~ x1, data = dat, family = "poisson",
                  coords = as.matrix(dat[, c("s1", "s2")]), cor.fn = "matern",
                  params.list = list(phi = c(3, 7, 10), nu = c(0.25, 0.5, 1.5),
                                     boundary = c(0.5, 0.6)),
                  n.samples = 1000,
                  loopd.controls = list(method = "CV", CV.K = 10, nMC = 1000),
                  parallel = TRUE, solver = "ECOS", verbose = TRUE)

# print(mod1$solver.status)
# print(mod1$run.time)

post_samps <- stackedSampler(mod1)
post_beta <- post_samps$beta
print(t(apply(post_beta, 1, function(x) quantile(x, c(0.025, 0.5, 0.975)))))

post_z <- post_samps$z
post_z_summ <- t(apply(post_z, 1, function(x) quantile(x, c(0.025, 0.5, 0.975)))))

z_combn <- data.frame(z = dat$z_true,
                    zL = post_z_summ[, 1],
                    zM = post_z_summ[, 2],
                    zU = post_z_summ[, 3])

library(ggplot2)
plot_z <- ggplot(data = z_combn, aes(x = z)) +
  geom_errorbar(aes(ymin = zL, ymax = zU),
               width = 0.05, alpha = 0.15,
               color = "skyblue") +
  geom_point(aes(y = zM), size = 0.25,
            color = "darkblue", alpha = 0.5) +
  geom_abline(slope = 1, intercept = 0,
             color = "red", linetype = "solid") +
  xlab("True z") + ylab("Posterior of z") +
  theme_bw() +
  theme(panel.background = element_blank(),
        aspect.ratio = 1)
```

spLMexact

*Univariate Bayesian spatial linear model***Description**

Fits a Bayesian spatial linear model with spatial process parameters and the noise-to-spatial variance ratio fixed to a value supplied by the user. The output contains posterior samples of the fixed effects, variance parameter, spatial random effects and, if required, leave-one-out predictive densities.

Usage

```
spLMexact(
  formula,
  data = parent.frame(),
  coords,
  cor.fn,
  priors,
  spParams,
  noise_sp_ratio,
  n.samples,
  loopd = FALSE,
  loopd.method = "exact",
  verbose = TRUE,
  ...
)
```

Arguments

formula	a symbolic description of the regression model to be fit. See example below.
data	an optional data frame containing the variables in the model. If not found in data, the variables are taken from environment(formula), typically the environment from which spLMexact is called.
coords	an $n \times 2$ matrix of the observation coordinates in \mathbb{R}^2 (e.g., easting and northing).
cor.fn	a quoted keyword that specifies the correlation function used to model the spatial dependence structure among the observations. Supported covariance model keywords are: 'exponential' and 'matern'. See below for details.
priors	a list with each tag corresponding to a parameter name and containing prior details.
spParams	fixed value of spatial process parameters.
noise_sp_ratio	noise-to-spatial variance ratio.
n.samples	number of posterior samples to be generated.
loopd	logical. If loopd=TRUE, returns leave-one-out predictive densities, using method as given by loopd.method. Default is FALSE.

loopd.method	character. Ignored if loopd=FALSE. If loopd=TRUE, valid inputs are 'exact' and 'PSIS'. The option 'exact' corresponds to exact leave-one-out predictive densities which requires computation almost equivalent to fitting the model n times. The option 'PSIS' is faster and finds approximate leave-one-out predictive densities using Pareto-smoothed importance sampling (Gelman <i>et al.</i> 2024).
verbose	logical. If verbose = TRUE, prints model description.
...	currently no additional argument.

Details

Suppose $\chi = (s_1, \dots, s_n)$ denotes the n spatial locations the response y is observed. With this function, we fit a conjugate Bayesian hierarchical spatial model

$$y \mid z, \beta, \sigma^2 \sim N(X\beta + z, \delta^2 \sigma^2 I_n), \quad z \mid \sigma^2 \sim N(0, \sigma^2 R(\chi; \phi, \nu)),$$

$$\beta \mid \sigma^2 \sim N(\mu_\beta, \sigma^2 V_\beta), \quad \sigma^2 \sim \text{IG}(a_\sigma, b_\sigma)$$

where we fix the spatial process parameters ϕ and ν , the noise-to-spatial variance ratio δ^2 and the hyperparameters μ_β , V_β , a_σ and b_σ . We utilize a composition sampling strategy to sample the model parameters from their joint posterior distribution which can be written as

$$p(\sigma^2, \beta, z \mid y) = p(\sigma^2 \mid y) \times p(\beta \mid \sigma^2, y) \times p(z \mid \beta, \sigma^2, y).$$

We proceed by first sampling σ^2 from its marginal posterior, then given the samples of σ^2 , we sample β and subsequently, we sample z conditioned on the posterior samples of β and σ^2 (Banerjee 2020).

Value

An object of class spLMexact, which is a list with the following tags -

samples a list of length 3, containing posterior samples of fixed effects (beta), variance parameter (sigmaSq), spatial effects (z).

loopd If loopd=TRUE, contains leave-one-out predictive densities.

model.params Values of the fixed parameters that includes phi (spatial decay), nu (spatial smoothness) and noise_sp_ratio (noise-to-spatial variance ratio).

The return object might include additional data used for subsequent prediction and/or model fit evaluation.

Author(s)

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References

- Banerjee S (2020). "Modeling massive spatial datasets using a conjugate Bayesian linear modeling framework." *Spatial Statistics*, **37**, 100417. ISSN 2211-6753. doi:[10.1016/j.spasta.2020.100417](https://doi.org/10.1016/j.spasta.2020.100417).
- Vehtari A, Simpson D, Gelman A, Yao Y, Gabry J (2024). "Pareto Smoothed Importance Sampling." *Journal of Machine Learning Research*, **25**(72), 1-58. URL <https://jmlr.org/papers/v25/19-556.html>.

See Also[spLMstack\(\)](#)**Examples**

```

# load data
data(simGaussian)
dat <- simGaussian[1:100, ]

# setup prior list
muBeta <- c(0, 0)
VBeta <- cbind(c(1.0, 0.0), c(0.0, 1.0))
sigmaSqIGa <- 2
sigmaSqIGb <- 0.1
prior_list <- list(beta.norm = list(muBeta, VBeta),
                  sigma.sq.ig = c(sigmaSqIGa, sigmaSqIGb))

# supply fixed values of model parameters
phi0 <- 3
nu0 <- 0.75
noise.sp.ratio <- 0.8

mod1 <- spLMexact(y ~ x1, data = dat,
                 coords = as.matrix(dat[, c("s1", "s2")]),
                 cor.fn = "matern",
                 priors = prior_list,
                 spParams = list(phi = phi0, nu = nu0),
                 noise_sp_ratio = noise.sp.ratio,
                 n.samples = 100,
                 loopd = TRUE, loopd.method = "exact")

beta.post <- mod1$samples$beta
z.post.median <- apply(mod1$samples$z, 1, median)
dat$z.post.median <- z.post.median
plot1 <- surfaceplot(dat, coords_name = c("s1", "s2"),
                    var_name = "z_true")
plot2 <- surfaceplot(dat, coords_name = c("s1", "s2"),
                    var_name = "z.post.median")

plot1
plot2

```

Description

Fits Bayesian spatial linear model on a collection of candidate models constructed based on some candidate values of some model parameters specified by the user and subsequently combines inference by stacking predictive densities. See Zhang, Tang and Banerjee (2024) for more details.

Usage

```

spLMstack(
  formula,
  data = parent.frame(),
  coords,
  cor.fn,
  priors,
  params.list,
  n.samples,
  loopd.method,
  parallel = FALSE,
  solver = "ECOS",
  verbose = TRUE,
  ...
)

```

Arguments

formula	a symbolic description of the regression model to be fit. See example below.
data	an optional data frame containing the variables in the model. If not found in data, the variables are taken from <code>environment(formula)</code> , typically the environment from which <code>spLMstack</code> is called.
coords	an $n \times 2$ matrix of the observation coordinates in \mathbb{R}^2 (e.g., easting and northing).
cor.fn	a quoted keyword that specifies the correlation function used to model the spatial dependence structure among the observations. Supported covariance model keywords are: 'exponential' and 'matern'. See below for details.
priors	a list with each tag corresponding to a parameter name and containing prior details. If not supplied, uses defaults.
params.list	a list containing candidate values of spatial process parameters for the <code>cor.fn</code> used, and, noise-to-spatial variance ratio.
n.samples	number of posterior samples to be generated.
loopd.method	character. Valid inputs are 'exact' and 'PSIS'. The option 'exact' corresponds to exact leave-one-out predictive densities. The option 'PSIS' is faster, as it finds approximate leave-one-out predictive densities using Pareto-smoothed importance sampling (Gelman <i>et al.</i> 2024).
parallel	logical. If <code>parallel=FALSE</code> , the parallelization plan, if set up by the user, is ignored. If <code>parallel=TRUE</code> , the function inherits the parallelization plan that is set by the user via the function <code>future::plan()</code> only. Depending on the parallel backend available, users may choose their own plan. More details are available at https://cran.R-project.org/package=future .
solver	(optional) Specifies the name of the solver that will be used to obtain optimal stacking weights for each candidate model. Default is "ECOS". Users can use other solvers supported by the <code>CVXR-package</code> package.
verbose	logical. If TRUE, prints model-specific optimal stacking weights.
...	currently no additional argument.

Details

Instead of assigning a prior on the process parameters ϕ and ν , noise-to-spatial variance ratio δ^2 , we consider a set of candidate models based on some candidate values of these parameters supplied by the user. Suppose the set of candidate models is $\mathcal{M} = \{M_1, \dots, M_G\}$. Then for each $g = 1, \dots, G$, we sample from the posterior distribution $p(\sigma^2, \beta, z \mid y, M_g)$ under the model M_g and find leave-one-out predictive densities $p(y_i \mid y_{-i}, M_g)$. Then we solve the optimization problem

$$\begin{aligned} \max_{w_1, \dots, w_G} \quad & \frac{1}{n} \sum_{i=1}^n \log \sum_{g=1}^G w_g p(y_i \mid y_{-i}, M_g) \\ \text{subject to} \quad & w_g \geq 0, \sum_{g=1}^G w_g = 1 \end{aligned}$$

to find the optimal stacking weights $\hat{w}_1, \dots, \hat{w}_G$.

Value

An object of class `spLMstack`, which is a list including the following tags -

`samples` a list of length equal to total number of candidate models with each entry corresponding to a list of length 3, containing posterior samples of fixed effects (`beta`), variance parameter (`sigmaSq`), spatial effects (`z`) for that model.

`loopd` a list of length equal to total number of candidate models with each entry containing leave-one-out predictive densities under that particular model.

`n.models` number of candidate models that are fit.

`candidate.models` a matrix with `n_model` rows with each row containing details of the model parameters and its optimal weight.

`stacking.weights` a numeric vector of length equal to the number of candidate models storing the optimal stacking weights.

`run.time` a `proc_time` object with runtime details.

`solver.status` solver status as returned by the optimization routine.

The return object might include additional data that is useful for subsequent prediction, model fit evaluation and other utilities.

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References

Vehtari A, Simpson D, Gelman A, Yao Y, Gabry J (2024). "Pareto Smoothed Importance Sampling." *Journal of Machine Learning Research*, **25**(72), 1-58. URL <https://jmlr.org/papers/v25/19-556.html>.

Zhang L, Tang W, Banerjee S (2024). "Bayesian Geostatistics Using Predictive Stacking." [doi:10.48550/arXiv.2304.12414](https://doi.org/10.48550/arXiv.2304.12414).

See Also

[spLMexact\(\)](#), [spGLMstack\(\)](#)

Examples

```
# load data and work with first 100 rows
data(simGaussian)
dat <- simGaussian[1:100, ]

# setup prior list
muBeta <- c(0, 0)
VBeta <- cbind(c(1.0, 0.0), c(0.0, 1.0))
sigmaSqIGa <- 2
sigmaSqIGb <- 2
prior_list <- list(beta.norm = list(muBeta, VBeta),
                  sigma.sq.ig = c(sigmaSqIGa, sigmaSqIGb))

mod1 <- spLMstack(y ~ x1, data = dat,
                 coords = as.matrix(dat[, c("s1", "s2")]),
                 cor.fn = "matern",
                 priors = prior_list,
                 params.list = list(phi = c(1.5, 3),
                                    nu = c(0.5, 1),
                                    noise_sp_ratio = c(1)),
                 n.samples = 1000, loopd.method = "exact",
                 parallel = FALSE, solver = "ECOS", verbose = TRUE)

post_samps <- stackedSampler(mod1)
post_beta <- post_samps$beta
print(t(apply(post_beta, 1, function(x) quantile(x, c(0.025, 0.5, 0.975))))))

post_z <- post_samps$z
post_z_summ <- t(apply(post_z, 1,
                    function(x) quantile(x, c(0.025, 0.5, 0.975))))

z_combn <- data.frame(z = dat$z_true,
                    zL = post_z_summ[, 1],
                    zM = post_z_summ[, 2],
                    zU = post_z_summ[, 3])

library(ggplot2)
plot1 <- ggplot(data = z_combn, aes(x = z)) +
  geom_point(aes(y = zM), size = 0.25,
            color = "darkblue", alpha = 0.5) +
  geom_errorbar(aes(ymin = zL, ymax = zU),
              width = 0.05, alpha = 0.15) +
  geom_abline(slope = 1, intercept = 0,
            color = "red", linetype = "solid") +
  xlab("True z") + ylab("Stacked posterior of z") +
  theme_bw() +
  theme(panel.background = element_blank(),
        aspect.ratio = 1)
```

stackedSampler *Sample from the stacked posterior distribution*

Description

A helper function to sample from the stacked posterior distribution to obtain final posterior samples that can be used for subsequent analysis. This function applies on outputs of functions `splMstack()` and `spGLMstack()`.

Usage

```
stackedSampler(mod_out, n.samples)
```

Arguments

<code>mod_out</code>	an object of class <code>splMstack</code> or <code>spGLMstack</code> .
<code>n.samples</code>	(optional) If missing, inherits the number of posterior samples from the original output. Otherwise, it specifies number of posterior samples to draw from the stacked posterior. If it exceeds the number of posterior draws used in the original function, then a warning is thrown and the samples are obtained by resampling. It is recommended, to run the original function with enough samples.

Details

After obtaining the optimal stacking weights $\hat{w}_1, \dots, \hat{w}_G$, posterior inference of quantities of interest subsequently proceed from the *stacked* posterior,

$$\tilde{p}(\cdot | y) = \sum_{g=1}^G \hat{w}_g p(\cdot | y, M_g),$$

where $\mathcal{M} = \{M_1, \dots, M_g\}$ is the collection of candidate models.

Value

An object of class `stacked_posterior`, which is a list that includes the following tags -

beta samples of the fixed effect from the stacked joint posterior.

z samples of the spatial random effects from the stacked joint posterior.

In case of model output of class `splMstack`, the list additionally contains `sigmaSq` which are the samples of the variance parameter from the stacked joint posterior of the spatial linear model. For model output of class `spGLMstack`, the list also contains `xi` which are the samples of the fine-scale variation term from the stacked joint posterior of the spatial generalized linear model.

Author(s)

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See Also

[spLMstack\(\)](#), [spGLMstack\(\)](#)

Examples

```
data(simGaussian)
dat <- simGaussian[1:100, ]

mod1 <- spLMstack(y ~ x1, data = dat,
                 coords = as.matrix(dat[, c("s1", "s2")]),
                 cor.fn = "matern",
                 params.list = list(phi = c(1.5, 3),
                                   nu = c(0.5, 1),
                                   noise_sp_ratio = c(1)),
                 n.samples = 1000, loopd.method = "exact",
                 parallel = FALSE, solver = "ECOS", verbose = TRUE)

print(mod1$solver.status)
print(mod1$run.time)

post_samps <- stackedSampler(mod1)
post_beta <- post_samps$beta
print(t(apply(post_beta, 1, function(x) quantile(x, c(0.025, 0.5, 0.975))))))
```

surfaceplot

Make a surface plot

Description

Make a surface plot

Usage

```
surfaceplot(tab, coords_name, var_name, h = 8, col.pal, mark_points = FALSE)
```

Arguments

tab	a data-frame containing spatial co-ordinates and the variable to plot
coords_name	name of the two columns that contains the co-ordinates of the points
var_name	name of the column containing the variable to be plotted
h	integer; (optional) controls smoothness of the spatial interpolation as appearing in the MBA: :mba.surf() function. Default is 8.
col.pal	Optional; color palette, preferably divergent, use colorRampPalette function from grDevices . Default is 'RdYIBu'.
mark_points	Logical; if TRUE, the input points are marked. Default is FALSE.

Value

a ggplot object containing the surface plot

Author(s)

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 Sudipto Banerjee sudipto@ucla.edu

Examples

```
data(simGaussian)
plot1 <- surfaceplot(simGaussian, coords_name = c("s1", "s2"),
                    var_name = "z_true")
plot1

# try your favourite color palette
col.br <- colorRampPalette(c("blue", "white", "red"))
col.br.pal <- col.br(100)
plot2 <- surfaceplot(simGaussian, coords_name = c("s1", "s2"),
                    var_name = "z_true", col.pal = col.br.pal)
plot2
```

 surfaceplot2

Make two side-by-side surface plots

Description

Make two side-by-side surface plots, particularly useful towards a comparative study of two spatial surfaces.

Usage

```
surfaceplot2(
  tab,
  coords_name,
  var1_name,
  var2_name,
  h = 8,
  col.pal,
  mark_points = FALSE
)
```

Arguments

tab	a data-frame containing spatial co-ordinates and the variables to plot
coords_name	name of the two columns that contains the co-ordinates of the points
var1_name	name of the column containing the first variable to be plotted
var2_name	name of the column containing the second variable to be plotted
h	integer; (optional) controls smoothness of the spatial interpolation as appearing in the MBA::mba.surf() function. Default is 8.

`col.pal` Optional; color palette, preferably divergent, use `colorRampPalette` function from `grDevices`. Default is `'RdYIBu'`.

`mark_points` Logical; if TRUE, the input points are marked. Default is FALSE.

Value

a list containing two ggplot objects

Author(s)

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Sudipto Banerjee sudipto@ucla.edu

Examples

```
data(simGaussian)
plots_2 <- surfaceplot2(simGaussian, coords_name = c("s1", "s2"),
                        var1_name = "z_true", var2_name = "y")
plots_2
```

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